

1. We have

$$\begin{aligned} P(X \in B) &\geq P((X \in B) \cap (X=Y)) = P((Y \in B) \cap (X=Y)) \\ &= 1 - P((Y \notin B) \cup (X \neq Y)) \geq P(Y \in B) - P(X \neq Y) \quad (1) \end{aligned}$$

Further

$$\begin{aligned} P(X \in B) &= P((X \in B) \cap (X=Y)) + P((X \in B) \cap (X \neq Y)) \\ &\leq P(Y \in B) + P(X \neq Y) \quad (2) \end{aligned}$$

Taken together

$$-P(X \neq Y) \stackrel{(1)}{\leq} P(X \in B) - P(Y \in B) \stackrel{(2)}{\leq} P(X \neq Y), \text{ as claimed}$$

2 (a)  $P(X_n \neq 0) = P(Y \leq \frac{1}{n}) = \frac{1}{n} \rightarrow 0$

$\therefore \lim_{n \rightarrow \infty} P(X_n \geq \epsilon) = 0$  for arbitrarily small  $\epsilon > 0$

This implies  $X_n \xrightarrow{P} 0$

Now let  $A_n = \{X_n \neq 0\}$ , then  $\{\lim_{n \rightarrow \infty} X_n \neq 0\} = \{\bigcap_{m \geq 1} \bigcup_{n \geq m} A_n\}$ ;  $A_n \supset A_{n+1}$  (all  $n$ )

$$P\left(\bigcap_{m=0}^{\infty} \bigcup_{n=m}^{\infty} A_n\right) = \lim_{m \rightarrow \infty} P\left(\bigcup_{n=m}^{\infty} A_n\right) = \lim_{m \rightarrow \infty} P(A_m) = 0 \quad \therefore X_n \xrightarrow{\text{a.s.}} 0$$

Now notice that  $P(X_n^2 = n) = \frac{1}{n}$ , so

$$E X_n^2 = n \cdot \frac{1}{n} = 1, \text{ thus } X_n \not\xrightarrow{\text{m.s.}} 0 \quad (X_n \text{ does not have an m.s. limit})$$

(b) Now  $X_n = \sqrt{n} \mathbb{1}_{\{Y_n \leq \frac{1}{n}\}}$

$$P(X_n \neq 0) = P(Y_n \leq \frac{1}{n}) \rightarrow 0, \text{ thus } X_n \xrightarrow{P} 0$$

To address a.s. convergence, consider independent events  $A_n = \{X_n \neq 0\}$ ;  $P(A_n) = \frac{1}{n}$  and  $\sum_n P(A_n) = \infty$

Borel-Cantelli tells us that  $P(A_n \text{ i.o.}) = 1$ , so

$$P\left(\lim_{n \rightarrow \infty} X_n = 0\right) = 0. \quad X_n \not\xrightarrow{\text{a.s.}} 0$$

Finally,  $E X_n^2 = n P(Y_n \leq \frac{1}{n}) = 1$ , so  $X_n \not\xrightarrow{\text{m.s.}} 0$

3. Consider the independent events  $X_n > \lambda_n$ ; if

$$\sum_{n=1}^{\infty} (1 - F(\lambda_n)) = \infty, \text{ then } P(\{X_n > \lambda_n\} \text{ i.o.}) = 1.$$

Since  $\{X_n > \lambda_n\} \subset A_n$ , in this case also  $\underline{P(A_n \text{ i.o.}) = 1}$

Now suppose that  $\sum_{n=1}^{\infty} (1 - F(\lambda_n)) < \infty$ , i.e.  $P(\{X_n > \lambda_n\} \text{ i.o.}) = 0$

Assume in addition that the sequence  $(\lambda_n)_{n \geq 1}$  is monotone increasing,  $\lambda_n \uparrow \infty$ .

We need to show that if  $X_n \leq \lambda_n$  a.s., also  $\max_{1 \leq m \leq n} X_m \leq \lambda_n$  a.s.

$1 = P(X_n > \lambda_n \text{ finitely many times})$

$$= P\left(\bigcap_{N \geq n} \{X_n > \lambda_N\} \text{ finitely many times}\right) \text{ (by monotonicity)}$$

$$= P\left(\bigcap_{n=1}^N \{X_n > \lambda_N\} \text{ finitely many times}\right) = P(A_N^c \text{ i.o.})$$

$-\ln(1 - (1-p)) = -\ln p$

4. (a)  $P_X(n) = p(1-p)^n$ ,  $n=0,1,\dots$

$$E\left(\frac{1}{1+X}\right) = p \sum_{n=0}^{\infty} \frac{(1-p)^n}{n+1} = \frac{p}{1-p} \sum_{n=1}^{\infty} \frac{(1-p)^n}{n} = \frac{p}{1-p} \ln \frac{1}{p}.$$

recognize the circled sum as  $-\ln(1 - (1-p)) = -\ln p$

$$(b) P(X_n \leq nt) = 1 - \left(1 - \frac{\lambda}{\lambda + n}\right)^{nt} = 1 - \left(1 - \frac{\lambda}{\lambda + n}\right)^{\frac{\lambda + n}{\lambda} \cdot \frac{\lambda}{\lambda + n} \cdot nt}$$

$$\xrightarrow{n \rightarrow \infty} \underline{1 - e^{-\lambda t}}, \text{ i.e., } \underline{\frac{X_n}{n} \xrightarrow{d} \text{Exp}(\lambda)}$$

CDF  $\text{Exp}(\lambda)$

$$5. (a) P(\min(X, Y) \leq n) = 1 - P(X > n \cap Y > n) = 1 - \left( \sum_{m=n+1}^{\infty} 2^{-m} \right)^2$$

$$= 1 - \left( \frac{2^{-n-1}}{1/2} \right)^2 = 1 - 4^{-n}$$

$$(b) P(X=Y) = \sum_{n=1}^{\infty} P(X=n)^2 = \sum_{n=1}^{\infty} 2^{-2n} = \frac{1/4}{3/4} = \frac{1}{3}$$

$$(c) P(X > Y) = \sum_{n=1}^{\infty} P(Y=n) \sum_{m=n+1}^{\infty} P(X=m) = \sum_{n=1}^{\infty} 2^{-n} \sum_{m=n+1}^{\infty} 2^{-m} = \sum_{n=1}^{\infty} 2^{-n} \frac{2^{-n-1}}{1/2}$$

$$= \sum_{n=1}^{\infty} 2^{-2n} = \frac{1}{3}$$

$$(d) P(X|Y) = \sum_{n=1}^{\infty} P(X=n) \sum_{m=1}^{\infty} P(Y=mn) = \sum_{n=1}^{\infty} 2^{-n} \sum_{m=1}^{\infty} (2^{-n})^m$$

$$= \sum_{n=1}^{\infty} 2^{-n} \frac{2^{-n}}{1-2^{-n}} = \sum_{n=1}^{\infty} \frac{1}{2^n(2^n-1)} \quad (\text{no closed-form expression : ( )})$$

$$(e) P(X \geq kY) = \sum_{n=1}^{\infty} P(Y=n) P(X \geq kn) = \sum_{n=1}^{\infty} 2^{-n} \sum_{m=kn}^{\infty} 2^{-m}$$

$$= \sum_{n=1}^{\infty} 2^{-n} \frac{2^{-kn}}{1-1/2} = 2 \sum_{n=1}^{\infty} 2^{-(k+1)n} = 2 \frac{2^{-(k+1)}}{1-2^{-(k+1)}} = \frac{2}{2^{k+1}-1}$$